Lecture 2 Conditional Probability and Bayes' Theorem

**Chao Song** 

College of Ecology Lanzhou University

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#### A motivating example

**Example:** A deck of cards (52 cards without jokers) is well shuffled and one card is drawn randomly. What is the probability that the card is a K?

$$P(K) = \frac{4}{52} = \frac{1}{13}$$

What is the probability of drawing a K if the card is known to be a face card (J, Q. K)?

$$P(K|J,Q,K) = \frac{4}{12} = \frac{1}{3}$$

#### Why do the two probabilities differ ?

Current knowledge (face card) has changed or restricted the sample space.

# **Conditional probability**

Given two events *A* and *B*. We denote the probability of event *A* happens given that event *B* is known to happen as P(A|B).

We can think of "given *B*" as specifying the new sample space for which, to determine P(A|B), we now want to calculate the probability of that part of A that is contained in B.



# **Conditional probability**

The conditional probability of event A given event B can be calculated as

$$P(A|B) = rac{P(AB)}{P(B)}$$

provided that P(B) > 0

Conditional probability satisfies the axioms for a probability function:

- $P(A|B) \ge 0$
- P(B|B) = 1
- if A<sub>1</sub>, A<sub>2</sub>, A<sub>3</sub>... are mutually exclusive events, then

 $P(A_1 \cup A_2 \cup \ldots \cup A_k|B) = P(A_1|B) + P(A_2|B) + \ldots + P(A_k|B)$ 

for each positive integer k, and

$$P(A_1 \cup A_2 \cup \ldots | B) = P(A_1 | B) + P(A_2 | B) + \ldots$$

for an infinite but countable number of events.

# **Conditional probability**

**Example**: A common test for AIDS is called ELISA. Among 1 million people who are given the test, we obtain results in the following table

	B <sub>1</sub> : AIDS	B <sub>2</sub> : No AIDS	Totals
A <sub>1</sub> : Positive	4885	73630	78515
A <sub>2</sub> : Negative	115	921370	921485
Totals	5000	995000	1000000

Find the following probabilities:  $P(B_1)$ ,  $P(A_1)$ ,  $P(A_1|B_2)$ ,  $P(B_1|A_1)$ .

## **General multiplication rule**

Intuitively, we can view events *A* and *B* both occur as a two step process: event *B* occurs and then event *A* occur given that event *B* has already occurred.

The formula for conditional probability can be used the other way around. Multiplying both side by P(B), we get the **general multiplication rule**:

P(AB) = P(B)P(A|B)

#### General multiplication rule for several events

The general multiplication rule can be extended to more than two events:

 $P(ABC) = P(A) \times P(B|A) \times P(C|A, B);$ 

 $P(ABCD) = P(A) \times P(B|A) \times P(C|A, B) \times P(D|A, B, C);$ 

 $P(ABCDE) = P(A) \times P(B|A) \times P(C|A, B) \times P(D|A, B, C) \times P(E | A, B, C, D).$ 

#### **General multiplication rule**

**Example**: A box contains 7 blue balls and 3 red balls. We randomly draw two balls successively without replacement. We want to compute the probability that the first draw is a red ball (*A*) and the second draw is a blue ball (*B*).

Using the definition of conditional probability, we have

$$P(AB) = P(A)P(B|A) = \frac{3}{10} \cdot \frac{7}{9} = \frac{7}{30}$$

Using the method of enumeration, we have

$$P(AB) = \frac{\mathbf{C}_3^1 \cdot \mathbf{C}_7^1}{\mathbf{A}_{10}^2} = \frac{7}{30}$$

**Comment**: We can compute probability by two seemingly different methods provided that our reasoning is consistent with the underlying assumptions.

Two events are **independent** if knowing the outcome of one provides no useful information about the outcome of the other.

- **Independent**: Knowing that the coin landed on a head on the first toss does not provide any information for determining what the coin will land in the second toss.
- **Dependent**: Knowing that the first card drawn from a deck is an ace provide information on determining the probability of drawing an ace in the second draw.

**Definition**: Events *A* and *B* are independent if and only if  $P(AB) = P(A) \times P(B)$ . In other words,  $P(AB) = P(A) \times P(B)$  is a necessary and sufficient condition for *A* and *B* being independent.

- This comes from the general multiplication rule. P(AB) = P(A) × P(B|A) in which P(B|A) reduces to P(B) when events A and B are independent.
- More generally, if events  $A_1, A_2, \ldots, A_k$  are independent,

$$P(A_1A_2...A_k) = P(A_1) \times P(A_2) \times ... P(A_k)$$

If events *A* and *B* are independent, events *B* and *C* are independent, are events *A* and *C* necessarily independent?

If two events *A* and *B* are disjoint with  $P(A) \neq 0$  and  $P(B) \neq 0$ , are they independent?

As estimated in 2012, of the US population,

- 13.4% were 65 or older, and
- 52% of the population were male.

True or False: 0.134  $\times$  0.52  $\approx$  7% of the US population were males aged 65 or older.

#### The answer is false

- Age and gender are not independent. On average women live longer than men. There are more old women than old men;
- According to survey data, among those 65 or older in the US, 44% are male, not 52%. Thus,  $0.134 \times 0.44 \approx 5.9\%$  were males aged 65 or older in the US in 2012.

#### Law of total probability

**Law of total probability**: If events  $A_1, A_2, ..., A_n$  are pairwise disjoint events,  $B \subset \bigcup_{k=1}^n A_k$ , then

$$P(B) = \sum_{k=1}^{n} P(A_k) P(B|A_k)$$

**Corollary 1**: If events  $A_1, A_2, ..., A_n$  is a partition of the sample space, i.e., a set of pairwise disjoint events whose union is the entire sample space, then

$$P(B) = \sum_{k=1}^{n} P(A_k) P(B|A_k)$$

**Corollary 2**: For any events *A* and *B*, the probability of *B* can be calculated as P(B) = P(A)P(B|A) + P(A')P(B|A').

# Law of total probability

**Seroprevalence adjustment**: China Center for Disease Control and Prevention conducted a serological survey in April, 2020 in Wuhan and found 4.43% positive test. The test has a sensitivity of 90% and specificity of 98%. Does that mean that 4.43% of the Wuhan population were once infected?

**Solution**: Let *A* denote a positive test and *B* denote a true infection. Using the law of total probability, we have

$$P(A) = P(B)P(A|B) + P(B')P(A|B').$$

• 
$$P(A) = 4.43\%;$$

- Sensitivity P(A|B) = 0.9;
- Specificity P(A'|B') = 0.98;

Plug these values to the equation above, we get

$$0.0443 = P(B) \times 0.9 + (1 - P(B)) \times (1 - 0.98)$$

Solving the equation above for P(B), we arrive at P(B) = 2.76%.

**Bayes' theorem**, named after Thomas Bayes, describes the probability of an event, based on prior knowledge of conditions that might be related to the event:

$$P(A|B) = rac{P(B|A)P(A)}{P(B)},$$

where A and B are events and  $P(B) \neq 0$ .



Thomas Bayes (1702-1761)

**Disease survey problem**: Wu et al. (2019) estimated that people living with human immunodeficiency virus (HIV) has risen to more than 1.25 million in China, roughly 0.09% of the total population. A blood test for HIV typically has 95% accuracy, i.e., the test correctly detects positive cases or negative case 95% times. If a person is tested positive, what is the probability that this person is infected with HIV?

**Solution**: Let *A* denote that a person has HIV and *B* denote a positive test. We know P(A) = 0.0009, P(B|A) = 0.95, and P(B'|A') = 0.95. We want to know P(A|B).

Using Bayes' theorem, we have

$$P(A|B) = \frac{P(A)P(B|A)}{P(B)} = \frac{0.0009 \times 0.95}{P(B)},$$

Using law of total probability, we have

$$P(B) = P(A)P(B|A) + P(A')P(B|A')$$
  
= 0.0009 × 0.95 + (1 - 0.0009) × (1 - 0.95)  
= 0.05081

Finally, we have P(A|B) = 1.68%

For a rare disease, if you get a positive test, the. chance that you indeed have that disease is very low.

**Disease survey problem**: Will a more accurate test help? If we improve the test sensitivity and specificity to 99.9%. What is the probability of having HIV if a person tested positive?

Solution: Using Bayes' theorem and the law of total probability:

$$P(A|B) = \frac{P(A)P(B|A)}{P(A)P(B|A) + P(A')P(B|A')}$$
$$= \frac{0.0009 \times 0.999}{0.0009 \times 0.999 + (1 - 0.0009) \times (1 - 0.999)}$$
$$= 47.36\%.$$

**Implications**: We should be cautious when confirming a rare disease, even if the diagnostic test is highly accurate.